

# BST 676 — Spring 2010 — Dr. Charnigo

## Written Assignment 2

Written Assignment 2 is due on Wednesday 17 February at the end of class. You are encouraged to work in groups of two or three, but you may work individually if you prefer.

[25] 1. This exercise will illustrate that the method of moments and maximum likelihood may not produce the same estimates.

[10] a. Find the method of moments estimate of  $\theta$  for Example #5 in Unit II.

[15] b. Find the maximum likelihood estimate of  $\theta$  for Example #3 in Unit II. [Let  $\zeta_1$  and  $\zeta_2$  denote the components of the two-dimensional vector  $\zeta$ . Differentiate the log likelihood in  $\zeta_1$  and set the result equal to zero. Then differentiate the log likelihood in  $\zeta_2$  and set the result equal to zero. The solutions to these two equations will be the components of  $\hat{\theta}$ .]

[25] 2. Let  $X_1, X_2, \dots, X_n$  be independent and identically distributed with probability density function  $f(x; \theta) := 3\theta^3/x^4$  for  $x \in [\theta, \infty)$  and  $\theta \in \Theta := (0, \infty)$ .

[10] a. Find the method of moments estimate of  $\theta$ .

[15] b. Find the maximum likelihood estimate of  $\theta$ .

[25] 3. Let  $X_1, X_2, \dots, X_n$  be independent and identically distributed with probability mass function  $f(x; \theta) := \theta^x(1 - \theta)^{1-x}$  for  $x \in \{0, 1\}$  and  $\theta \in \Theta := [0, 1]$ . [If  $\theta = 0$  or  $\theta = 1$ , then the distribution is degenerate.]

[10] a. Find the maximum likelihood estimate of  $\theta$ .

[15] b. Impose prior distribution  $p(\theta) \propto \theta^a(1 - \theta)^b$  for specified nonnegative constants  $a$  and  $b$ . Find the posterior mean and posterior mode estimates of  $\theta$ .

[25] 4. In their 2008 *Journal of Applied Statistics* paper “Omnibus testing and gene filtration in microarray data analysis”, Dai and Charnigo define a test statistic  $D_n$  such that, when the null hypothesis is true,  $nD_n \xrightarrow{L} \tau_1 Y_1 + \tau_2 Y_2$ , where  $\tau_1, \tau_2$  are known positive constants and  $Y_1, Y_2$  are independent chi-square random variables on one degree of freedom. If  $nD_n$  is larger than the 0.95 quantile of  $\tau_1 Y_1 + \tau_2 Y_2$ , we reject the null hypothesis at approximate significance level 0.05.

[10] a. Suppose that we wish to approximate  $\tau_1 Y_1 + \tau_2 Y_2$  by  $aY$ , where  $a$  is a nonnegative constant and  $Y$  is a chi-square random variable on  $\nu$  degrees of freedom. Propose choices for  $a$  and  $\nu$ .

[15] b. Suppose further that  $\tau_1 = 1/(4\sqrt{\pi})$  and  $\tau_2 = 3/(16\sqrt{\pi})$ . Let  $\chi_{\nu,0.95}^2$  denote the 0.95 quantile of the chi-square distribution on  $\nu$  degrees of freedom, which you can extract from R or other software of your choice. Perform a simulation experiment with R or other software of your choice to estimate  $P(\tau_1 Y_1 + \tau_2 Y_2 > a\chi_{\nu,0.95}^2)$ . [Simulate 10000 realizations of  $\tau_1 Y_1 + \tau_2 Y_2$  and record how many of them exceed  $a\chi_{\nu,0.95}^2$ .] Comment on the implications of using  $aY$  as an approximation to  $\tau_1 Y_1 + \tau_2 Y_2$  in the context of the above hypothesis test.