

# STA 623 – Fall 2011 – Dr. Charnigo

## Section 1.3: Conditional Probability and Independence

*Definition of conditional probability.* Let  $A$  and  $B$  be events that belong to the sigma field. (Hereafter all events mentioned will belong to the sigma field unless explicitly stated otherwise.) If  $P(B) > 0$ , then we define the conditional probability of  $A$  given  $B$  as

$$P(A|B) := P(A \cap B)/P(B).$$

In this context, we sometimes refer to  $P(A)$  as an unconditional probability. Intuitively,  $P(A|B)$  is an updated version of  $P(A)$  given the knowledge that event  $B$  has occurred.

**Example (definition of conditional probability).** Let  $A$  denote the event that a randomly selected person smokes, and let  $B$  denote the event that a randomly selected person develops lung cancer. Suppose that  $P(B) = 0.02$ ,  $P(A) = 0.25$ , and  $P(B|A) = 0.05$ . Then  $P(A \cap B) =$  \_\_\_\_\_, so that  $P(A|B) =$  \_\_\_\_\_. Moreover,  $P(A^c \cap B) =$  \_\_\_\_\_, so that  $P(A^c|B) =$  \_\_\_\_\_. Is the last answer intuitively obvious?

*Useful results for conditional probabilities.* Assuming that the axioms for unconditional probabilities are satisfied and that  $P(D) > 0$ , we have the following useful results for conditional probabilities.

1.  $P(A|D) \geq 0$ .
2.  $P(S|D) = 1$ .
3.  $P(\cup_{i=1}^{\infty} A_i|D) = \sum_{i=1}^{\infty} P(A_i|D)$  whenever  $A_i \cap A_j = \emptyset$  for  $i \neq j$ .
4.  $P(\emptyset|D) = 0$ .

5.  $P(A|D) \leq 1$ .
6.  $P(A^c|D) = 1 - P(A|D)$ .
7.  $P(B \cap A^c|D) = P(B|D) - P(B \cap A|D)$ .
8.  $P(A \cup B|D) = P(A|D) + P(B|D) - P(A \cap B|D)$ .
9. If  $A \subset B$ , then  $P(A|D) \leq P(B|D)$ .
10.  $P(\cup_{i=1}^{\infty} C_i|D) \leq \sum_{i=1}^{\infty} P(C_i|D)$ .
11. If  $C_1, C_2, \dots$  is a partition, then  $P(A|D) = \sum_{i=1}^{\infty} P(A \cap C_i|D)$ .

**Example (useful results for conditional probabilities).** To verify result 3, we write

$$P(\cup_{i=1}^{\infty} A_i|D) = P(\cup_{i=1}^{\infty} (A_i \cap D))/P(D) = \sum_{i=1}^{\infty} P(A_i|D).$$

The justification for the second equality is that, if  $A_i \cap A_j = \emptyset$ , then

$$(A_i \cap D) \cap (A_j \cap D) = (A_i \cap A_j) \cap D = \emptyset \cap D = \emptyset.$$

To verify result 9, we note that  $A \cap D \subset B \cap D$ . (How do we know this?) Then  $P(A \cap D) \leq P(B \cap D)$ , whence

$$P(A|D) = P(A \cap D)/P(D) \leq P(B \cap D)/P(D) = P(B|D).$$

*Iterating conditional probabilities.* Suppose that  $P(B \cap C) > 0$ . Then

$$P(A \cap B \cap C) = P(A|B \cap C)P(B \cap C) = P(A|B \cap C)P(B|C)P(C).$$

Moreover, since the left side is also equal to  $P(A \cap B|C)P(C)$ , we see that

$$P(A \cap B|C) = P(A|B \cap C)P(B|C).$$

(How do we know that  $P(C) > 0$ ?) The above formulas may be useful in situations when the conditional probabilities  $P(A|B \cap C)$  and  $P(B|C)$  are intuitively obvious while  $P(A \cap B|C)$  and  $P(A \cap B \cap C)$  are less readily apparent.

**Example (iterating conditional probabilities).** I am dealt three cards (without replacement) from a well-shuffled standard 52-card deck. Let  $C$  be the event that the suit of the first card is diamonds,  $B$  be the event that the suit of the second card is diamonds, and  $A$  be the event that the suit of the third card is diamonds. Then  $P(C) = 13/52$ ,  $P(B|C) = \frac{12}{51}$ , and  $P(A|B \cap C) = \frac{11}{50}$ . So  $P(A \cap B|C) = \frac{12}{51} \cdot \frac{11}{50}$  and  $P(A \cap B \cap C) = \frac{13}{52} \cdot \frac{12}{51} \cdot \frac{11}{50}$ .

*Bayes' Theorem.* Let  $A_1, A_2, \dots$  be a partition of  $S$  such that  $P(A_i) > 0$  for  $i \in \{1, 2, \dots\}$ . Then for any event  $B$  and any  $i \in \{1, 2, \dots\}$  we have

$$P(A_i|B) = \frac{P(B|A_i)P(A_i)}{\sum_{j=1}^{\infty} P(B|A_j)P(A_j)}.$$

A similar result also holds for a finite partition  $A_1, A_2, \dots, A_k$ ,

$$P(A_i|B) = \frac{P(B|A_i)P(A_i)}{\sum_{j=1}^k P(B|A_j)P(A_j)}.$$

In particular, with  $k = 2$  we have (upon a minor change in notation)

$$P(A|B) = \frac{P(B|A)P(A)}{P(B|A)P(A) + P(B|A^c)P(A^c)}.$$

This last version of Bayes' Theorem provides a recipe for computing  $P(A|B)$  if we have available  $P(B|A)$  (and two other probabilities — why two rather than three?).

**Example (Bayes' Theorem).** Bayes' Theorem is obvious from the definition of conditional probability if we can verify that the denominator is  $P(B)$ . (Equality of the denominator to  $P(B)$  is sometimes called the Law of Total Probability.) Suggestions?

Again, let  $A$  denote the event that a randomly selected person smokes, and let  $B$  denote the event that a randomly selected person develops lung cancer. As before, suppose that  $P(A) = 0.25$  and  $P(B|A) = 0.05$ . But now suppose that we are not given  $P(B)$ . Rather, suppose that we are told that 99% of non-smokers do not develop lung cancer. How can we find  $P(A|B)$  with just the information provided here?

*Independence of two events.* Suppose that  $A$  and  $B$  are events such that  $0 < P(A) < 1$ ,  $0 < P(B) < 1$ , and  $P(A|B) = P(A)$ . Thus, knowing that  $B$  has occurred does not lead us to revise the probability that  $A$  will occur. In this case, multiplying both sides of  $P(A|B) = P(A)$  by  $P(B)$  yields

$$P(A \cap B) = P(A|B)P(B) = P(A)P(B).$$

Moreover, since  $P(A \cap B)$  can also be expressed as  $P(B|A)P(A)$ , we see that  $P(B|A) = P(B)$ . If any one of these three equivalent conditions holds —  $P(A|B) = P(A) = P(A|B^c)$ ,  $P(A \cap B) = P(A)P(B)$ , or  $P(B|A) = P(B) = P(B|A^c)$  — we say that  $A$  and  $B$  are independent. (How are we able to go from  $P(A|B) = P(A)$  to  $P(A|B) = P(A) = P(A|B^c)$  and from  $P(B|A) = P(B)$  to  $P(B|A) = P(B) = P(B|A^c)$ ?)

If no restrictions on  $P(A)$  or  $P(B)$  are made, so we are not sure whether  $P(A|B)$  and  $P(B|A)$  are defined, then we characterize independence by the equation  $P(A \cap B) = P(A)P(B)$ .

If  $A$  and  $B$  are independent, then so are  $A^c$  and  $B^c$ ,  $A^c$  and  $B$ , and  $A$  and  $B^c$ .

**Example (independence of two events).** Let  $A$  and  $B$  be two events. If  $P(A) = 0$ , then  $A$  and  $B$  are independent. (How do we know this?) If  $P(B) = 1$ , then  $A$  and  $B$  are independent. (How do we know this?)

To verify that independence of  $A$  and  $B$  implies independence of  $A^c$  and  $B$ , we write

$$P(A^c \cap B) = P(B) - P(A \cap B) = P(B) - P(A)P(B) = P(B)P(A^c).$$

We can similarly prove that independence of  $A$  and  $B$  implies independence of  $A^c$  and  $B^c$  and of  $A$  and  $B^c$ .

Suppose once more that I am dealt three cards (without replacement) from a well-shuffled standard 52-card deck. Let  $C$  be the event that the suit of the first card is diamonds,  $B$  be the event that the suit of the second card is diamonds, and  $A$  be the event that the suit of the third card is diamonds. Are  $C$  and  $B$  independent?

*Independence of three or more events.* Let  $A_1, A_2, \dots, A_n$  be events such that for any subcollection  $A_{i_1}, A_{i_2}, \dots, A_{i_k}$  we have

$$P(\cap_{j=1}^k A_{i_j}) = \prod_{j=1}^k P(A_{i_j}).$$

Then we say that  $A_1, A_2, \dots, A_n$  are independent. Independence of  $n$  events thus entails not only

$$P(A_1 \cap A_2 \cap \dots \cap A_n) = P(A_1)P(A_2) \dots P(A_n)$$

but also

$$P(A_1 \cap A_2) = P(A_1)P(A_2),$$

$$P(A_1 \cap A_2 \cap A_3) = P(A_1)P(A_2)P(A_3),$$

$$P(A_4 \cap A_{18}) = P(A_4)P(A_{18}),$$

and many other equalities.